# A Lower Bound for the Memory Capacity in the Potts-Hopfield Model 

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#### Abstract

We consider Potts-Hopfield networks of size $N$. We prove the result: $\exists \alpha_{c}>0$ such that for all $0<\alpha<\alpha_{c}$ we can find $\delta, \varepsilon>0$ in such a way that, when $N \rightarrow \infty$, we can store $\alpha N$ patterns, all of them being sorrounded by $\varepsilon$-energy barriers at distance $\delta$.


KEY WORDS: Neural network; Hopfield model; Potts model.

## INTRODUCTION

Potts-Hopfield networks correspond to $q$-state neural networks. They have been recently used to model some smoothing techniques and coloring rules (see, for instance, ref. 6). We consider the problem of the memory capacity. We remark that we are storring colored patterns with $q$ colors instead of black and white patterns as would be the case in the Hopfield model.

In our main theorem we consider the relation between local minima of the energy function and the stored patterns. We prove an analog of the result shown by Newman for the Hopfield model. ${ }^{(1)}$ To describe it, let us consider a network of $q$-state neurons, of size $N$. We prove that there exists $\alpha_{c}>0$ such that for any $0<\alpha<\alpha_{c}$ we can find $\varepsilon>0$ and $0<\delta<1 / 2$ such that, in the limit $N \rightarrow \infty$, we can store $\alpha N$ patterns, all of them being sorrounded by $\varepsilon$-energy barriers at distance $\delta$.

[^0]Our proof follows that of Newman. But some interesting technical points must be solved.

In Section 1 we describe the model and we show that it reduces to the Hopfield networks when $q=2$. We also code the Hamiltonian as well as the dynamics with a 2 -state neural network and at the end of this section we state the main result. The proofs are in Section 2.

A complete discussion of Potts model can be found in ref. 7.

## 1. DESCRIPTION OF THE MODEL AND MAIN RESULT

The model has $N$ neurons whose configurations are described by an $N$-vector $\sigma=\left(\sigma_{1}, \ldots, \sigma_{N}\right)$, where for any $i \in\{1, \ldots, N\}, \sigma_{i} \in\{0,1, \ldots, q-1\}=Q$, and $q$ is a positive integer $\geqslant 2$.

There are $m \in \mathbb{N}$ patterns $\left(\xi^{\mu}\right)_{\mu=1}^{m}$ to be stored, each of which is an $N$-vector $\xi^{\mu}=\left(\xi_{1}^{\mu}, \ldots, \xi_{N}^{\mu}\right)$, where $\forall \mu \in\{1, \ldots, m\}, \forall i \in\{1, \ldots, N\}, \xi_{i}^{\mu} \in Q$. We consider the random case, which corresponds to choosing $\left(\xi_{i}^{\mu}\right)$ i.i.d. with $\mathbb{P}\left(\tilde{\xi}_{i}^{\mu}=s\right)=1 / q \forall s \in Q$.

Given the patterns $\xi^{\mu}$, we define the energy

$$
\begin{equation*}
H(\sigma)=-\sum_{\mu=1}^{m}\left[\sum_{i=1}^{N}\left\{\delta\left(\xi_{i}^{\mu}, \sigma_{i}\right)-\frac{1}{q}\right\}\right]^{2} \quad \text { for } \quad \sigma \in Q^{N} \tag{1.1}
\end{equation*}
$$

where $\delta(\cdot, \cdot)$ is the Kronecker symbol.
Let us associate a possible sequential dynamics compatible with the Hamiltonian. Assume we have updated the $i$ th neuron of the configuration $\sigma=\left(\sigma_{1}, \ldots, \sigma_{N}\right)$ and call $\sigma^{\prime}=\left(\sigma_{1}^{\prime}, \ldots, \sigma_{N}^{\prime}\right)$ the new configuration. So

$$
\sigma_{j}^{\prime}=\sigma_{j} \quad \text { for any } \quad j \neq i
$$

It can be easily shown that

$$
\begin{aligned}
H\left(\sigma^{\prime}\right)-H(\sigma)= & -2 \sum_{\mu=1}^{m} \delta\left(\xi_{i}^{\mu}, \sigma_{i}^{\prime}\right)\left\{\left[\sum_{\substack{j=1 \\
j \neq i}}^{N}\left(\delta\left(\xi_{j}^{\mu}, \sigma_{j}\right)-\frac{1}{q}\right)\right]+\frac{1}{2}-\frac{1}{q}\right\} \\
& \left.+2 \sum_{\mu=1}^{m} \delta\left(\xi_{i}^{\mu}, \sigma_{i}\right)\left\{\sum_{\substack{j=1 \\
j \neq i}}^{N}\left(\delta\left(\xi_{j}^{\mu}, \sigma_{j}\right)-\frac{1}{q}\right)\right]-\frac{1}{q}\right\}
\end{aligned}
$$

Now define $\sigma_{i}^{\prime}=s$, where $s$ is any one of the elements $s^{\prime}$ of $Q$, which is a maximizer in the following formula:

$$
\begin{equation*}
\max _{s^{\prime} \in Q} \sum_{\mu=1}^{m} \delta\left(\xi_{i}^{\mu}, s^{\prime}\right)\left\{\left[\sum_{\substack{j=1 \\ j \neq i}}^{N}\left(\delta\left(\xi_{j}^{\mu}, \sigma_{j}\right)-\frac{1}{q}\right)\right]+\frac{1}{2}-\frac{1}{q}\right\} \tag{1.2}
\end{equation*}
$$

Hence $H$ is a Lyapunov function of this dynamics, that is, it decreases with the evolution. In particular, starting with some initial configuration $\sigma(0)$, the trajectory will follow the decreasing part of the graph of the function $\sigma \rightarrow H(\sigma)$, and will eventually reach a local minimum, where it will stop.

To analyze the expressions (1.1) and (1.2) in the case $q=2$, transform the $\{0,1\}$ variables $\sigma_{i}, \xi_{i}^{\mu}$ into the $\{-1,1\}$ variables $\tilde{\sigma}_{i}=2 \sigma_{i}-1$, $\widetilde{\xi}_{i}^{\mu}=2 \xi_{i}^{\mu}-1$. It is easily shown that $H(\sigma)$ given by (1.1) becomes

$$
H(\tilde{\sigma})=-\frac{1}{4} \sum_{\mu=i}^{N} \sum_{i=1}^{N} \sum_{j=1}^{N} \xi_{i} \xi_{j}^{\mu} \tilde{\sigma}_{i} \tilde{\sigma}_{j}
$$

which is the usual Lyapunov functional. ${ }^{(1)}$ Now the dynamics (1.2) is written into the new variables as

$$
\tilde{\sigma}_{i}=\operatorname{sign}\left\{\sum_{i=1}^{N} \sum_{\substack{j=1 \\ j \neq i}}^{N} \tilde{\xi}_{i}^{\mu} \tilde{\xi}_{j}^{\mu} \tilde{\sigma}_{j}\right\}
$$

which is the sequential dynamics of the Hopfield model. ${ }^{(1)}$
Furthermore, the general Potts-Hopfield model given by Hamiltonian (1.1) can be coded like a Hopfield network of size $q N$, the states of each neuron being $\{0,1\}$. In this purpose, associate to $\xi^{\mu}$ and $\sigma$ the following $q N$-vectors:

$$
\xi_{i, s}^{\mu}=\delta\left(\xi_{i}^{\mu}, s\right), \quad \hat{\sigma}_{i, s}=\delta\left(\sigma_{i}, s\right)
$$

Hence the Hamiltonian (1.1) becomes

$$
\begin{aligned}
H(\hat{\sigma})= & -\sum_{\mu=1}^{m} \sum_{i=1}^{N} \sum_{s=1}^{q} \sum_{j=1}^{N} \sum_{u=1}^{q} \hat{\xi}_{i, s}^{\mu} \hat{\xi}_{j, u}^{\mu} \hat{\sigma}_{i, s} \hat{\sigma}_{j, u} \\
& +2 N \sum_{\mu=1}^{m} \sum_{i=1}^{N} \sum_{s=1}^{q} \xi_{i, s}^{\mu} \hat{\sigma}_{i, s}-m N^{2}
\end{aligned}
$$

This corresponds to a Hamiltonian associated to the following sequential dynamics, ${ }^{(6)}$ which we write for the updating of the $(i, s)$ neuron:

$$
\left(\hat{\sigma}_{i s}\right)^{\prime}=1\left(\sum_{\mu=1}^{m} \sum_{j=1}^{N} \sum_{u=1}^{q} \xi_{i, s}^{\mu} \xi_{j, u}^{\mu} \hat{\sigma}_{j, u}-N \xi_{i, s}^{\mu}\right)
$$

where $1(v)$ is the threshold function, $1(v)=1$ if $v \geqslant 0,=0$ if $v<0$. This is similar to the iteration of a $\{0,1\}$ Hopfield network with thresholds $N \xi_{i, s}^{\mu}$. The difference from the Hopfield model is that the set of variables $\left\{\xi_{i, s}^{\mu}: \mu=1, \ldots, m ; i=1, \ldots, N ; s=1, \ldots, q\right\}$ are not independent; in fact, the
relations $\sum_{s=1}^{q} \xi_{i, s}^{\mu}=1$ must be verified for each pair ( $i, \mu$ ). Then the statistical properties of the Potts-Hopfield model are not simple extensions of the Hopfield case.

We want to study the relation between the local minima of the energy $H(\sigma)$ and the stored patterns $\left\{\xi^{\mu}: \mu=1, \ldots, m\right\}$ choosen at random with uniform distribution.

In the Hopfield case ( $q=2$ ), if $m=N /(c \log N)$ for some real number $c$, it is proved that the original patterns are local minima. ${ }^{(4,5)}$ The case $m=\alpha N$ was studied by Amit et al. ${ }^{(3)}$ using a nonrigorous argument that strongly suggests the presence of a threshold $\alpha_{c}$ such that if $0<\alpha<\alpha_{c}\left(\alpha_{c} \sim 0.15\right)$, the Hamiltonian has minima that are near the patterns, while if $\alpha>\alpha_{c}$, this is not the case.

An important rigorous result in this direction was made by Newman, ${ }^{(2)}$ who proved that $\exists \alpha>0$ such that if $\alpha<\alpha$, then for almost all realizations of random patterns there exists, around each pattern, an energy barrier. Let us be more precise.

Call $d\left(\sigma, \sigma^{\prime}\right)$ the Hamming distance between two configurations:

$$
\begin{equation*}
d\left(\sigma, \sigma^{\prime}\right)=\sum_{i=1}^{N}\left[1-\delta\left(\sigma_{i}, \sigma_{i}^{\prime}\right)\right] \tag{1.3}
\end{equation*}
$$

that is, the number of sites where the two configurations disagree. For any configuration $\sigma$ and real number $\delta \in[0,1]$, denote by

$$
S(\sigma, \delta)=\left\{\sigma^{\prime}: d\left(\sigma, \sigma^{\prime}\right)=[\delta N]\right\}
$$

the sphere of radius [ $\delta N$ ].
Call $h(\sigma, \delta)=\min \left\{H\left(\sigma^{\prime}\right): \sigma^{\prime} \in S(\sigma, \delta)\right\}$. We shall say that there exists an energy barrier at distance $\delta$ around a pattern $\xi^{\mu}$ if for some $\varepsilon>0$

$$
\begin{equation*}
H\left(\xi^{\mu}\right)+\varepsilon N^{2}<\underline{h}\left(\xi^{\mu}, \delta\right) \tag{1.4}
\end{equation*}
$$

We remark that the presence of an energy barrier at distance $\delta$ around $\xi^{\mu}$ does not imply that $\xi^{\mu}$ is a local minimum, but merely that at least one minimum is inside a ball of radius [ $\delta N$ ] centered at $\xi^{\mu}$. If we accept the idea of storing patterns that could be retrievied with errors, the notion of energy barriers is a good one.

The main result of this paper is the extension to $q$-state neural networks of the Newman result for $q=2$.

Theorem. Given $q \geqslant 2$, there exists an $\underline{\alpha}_{c}=\underline{\alpha}_{c}(q)>0$ such that for any $\alpha \leqslant \underline{\alpha}_{c}$, we can find $\delta \in(0,1 / 2)$ and an $\varepsilon>0$ such that, in the limit
$N \rightarrow \infty$, for almost all configurations of patterns, all patterns are surrounded by energy barriers at distance $\delta$.

Precisely,

$$
\begin{equation*}
\mathbb{P}\left(\bigcup_{K=1}^{\infty} \bigcap_{N \geqslant K}\left\{\forall \mu \in\{1, \ldots, \alpha N\} ; \forall \sigma \in S\left(\xi^{\mu}, \delta\right) ; H(\sigma)>H\left(\xi^{\mu}\right)+\varepsilon N^{2}\right\}\right)=1 \tag{1.5}
\end{equation*}
$$

## 2. PROOF OF THE THEOREM

Let us first remark that, given $\mu \in\{1, \ldots, \alpha N\}$ and a pattern $\xi^{\mu}$, the set of configurations $\sigma$ such that $d\left(\xi^{\mu}, \sigma\right)=[\delta N]$ can be obtained by choosing first the subset $J \subset\{1, \ldots, N\}$, where $\sigma$ and $\xi^{\mu}$ disagree, and then $\sigma_{j} \neq \xi_{j}^{\mu}$ $\forall j \in J$. In particular,

$$
\begin{equation*}
\left|S\left(\xi^{\mu}, \delta\right)\right|=\binom{N}{[\delta N]}(q-1)^{[\delta N]} \tag{2.1}
\end{equation*}
$$

From (2.1) we get

$$
\begin{align*}
& \mathbb{P}\left(\exists \mu \subset\{1, \ldots, \alpha N\}, \exists \sigma \in S\left(\xi^{\mu}, \delta\right) ; H(\sigma) \leqslant H\left(\xi^{\mu}\right)+\varepsilon N^{2}\right) \\
& \quad \leqslant \alpha N\binom{N}{[\delta N]}(q-1)^{[\delta N]} \mathbb{P}\left(H\left(\xi_{J}^{\mu}\right)-H\left(\xi^{\mu}\right) \leqslant \varepsilon N^{2}\right) \tag{2.2}
\end{align*}
$$

where $\xi_{J}^{\mu}$ is any pattern in $S\left(\xi^{\mu}, \delta\right)$; one can choose $J=\{1,2, \ldots,[\delta N]\}$ and $\mu$ is any index $\in\{1, \ldots, \alpha N\}$.

In order to estimate the probability in the right-hand side of (2.2), we introduce some notations.

Let

$$
\begin{aligned}
V_{1} & =2 \frac{\delta}{q} \\
V_{2} & =4(1-\delta)\left(\frac{1}{q}\right)\left(1-\frac{1}{q}\right)+2 \delta\left(\frac{1}{q}\right)\left(1-\frac{2}{q}\right) \\
\gamma & =\delta\left[(1-\delta)+\left(1-\frac{2}{q}\right)\right]
\end{aligned}
$$

Define the variables

$$
\begin{equation*}
Z_{1, J}^{\mu^{\prime}}=\sum_{i \in J} \frac{\delta\left(\xi_{i}^{\mu^{\prime}},\left(\xi_{J}^{\mu}\right)_{i}\right)-\delta\left(\xi_{i}^{\mu^{\prime}}, \xi_{i}^{\mu}\right)}{\left(N V_{1}\right)^{1 / 2}} \tag{2.3}
\end{equation*}
$$

$$
\begin{align*}
Z_{2, J}^{\mu^{\prime}} & =\sum_{i \in J} \frac{\delta\left(\xi_{i}^{\mu^{\prime}},\left(\xi_{J}^{\mu}\right)_{i}\right)+\delta\left(\xi_{i}^{\mu^{\prime}}, \xi_{i}^{\mu}\right)-2 / q}{\left(N V_{2}\right)^{1 / 2}}  \tag{2.4}\\
Z_{2, J^{c}}^{\mu^{\prime}} & =\sum_{i \in J^{c}} \frac{2\left(\delta\left(\xi_{i}^{\mu^{\prime}}, \xi_{i}^{\mu}\right)-1 / q\right)}{\left(N V_{2}\right)^{1 / 2}} \tag{2.5}
\end{align*}
$$

Using the identity $x^{2}-y^{2}=(x-y)(x+y)$, it is not difficult to check that

$$
\begin{equation*}
H\left(\xi_{J}^{\mu}\right)-H\left(\xi^{\mu}\right)=-\left(V_{1} V_{2}\right)^{1 / 2} N \sum_{\substack{\mu^{\prime}=1 \\ \mu^{\prime} \neq \mu}}^{\alpha N} Z_{1, J}^{\mu^{\prime}}\left(Z_{2, J}^{\mu^{\prime}}+Z_{2, c^{c}}^{\mu^{\prime}}\right)+N^{2} \gamma \tag{2.6}
\end{equation*}
$$

In particular, the event which appears in the right-hand side of (2.2) is

$$
\begin{equation*}
B_{N} \equiv\left\{u \leqslant \frac{1}{\alpha N} \sum_{\substack{\mu^{\prime}=1 \\ \mu^{\prime} \neq \mu}}^{\alpha N} Z_{1, J}^{\mu^{\prime}}\left(Z_{2, J}^{\mu^{\prime}}+Z_{2, J^{c}}^{\mu^{\prime}}\right)\right\} \tag{2.7}
\end{equation*}
$$

where $u=(\gamma-\varepsilon) / \alpha\left(V_{1} V_{2}\right)^{1 / 2}$.
Now given $\mu$ and a fixed realization of $\xi^{\mu}$, it is easy to check that the family

$$
\left\{Z_{1, J}^{\mu^{\prime}}\left(Z_{2, J}^{\mu^{\prime}}+Z_{2, J^{c}}^{\mu^{\prime}}\right)\right\}_{\substack{\mu^{\prime} \\ \mu^{\prime} \neq \mu}}^{\alpha N}
$$

is a family of i.i.d. random variables.

Remark. We emphasize that, given $\xi^{\mu}$, the two random variables $\left(Z_{1, J}^{\mu^{\prime}}, Z_{2, j}^{\mu^{\prime}}\right)$ are not independent, but $Z_{1, J}^{\mu^{\prime}}$ and $Z_{2, J}^{\mu^{\prime}}$ are independent of $Z_{2, J^{c}}^{\mu}$. This phenomenon occurs only if $q>2$; if $q=2$, it can be verified that $Z_{2, J}^{\mu^{\prime}}=0 \forall \mu^{\prime} \in\{1, \ldots, \alpha N\}$. This is the main difference between the two cases.

It is easy to show that $\mathbb{E}\left(Z_{1, J}^{\mu^{\prime}} Z_{2, J}^{\mu^{\prime}}\right)=0$ [in fact, it follows from expression (2.11) below]; then $B_{N}$ is the event of a large deviation from the mean; it is natural to use the exponential Markov inequality to get

$$
\begin{equation*}
\mathbb{P}\left(B_{N}\right) \leqslant\left[\inf _{t \geqslant 0} e^{-t u} \mathbb{E}\left(e^{i Z_{1, J}\left(Z_{2, J}+Z_{2, J c}\right)}\right)\right]^{\alpha N} \tag{2.8}
\end{equation*}
$$

where $\left(Z_{1, J} ; Z_{2, J} ; Z_{2, J^{c}}\right)$ have the same distribution as any given $\left(Z_{1, J}^{\mu^{\prime}} ; Z_{2, J}^{\mu^{\prime}} ; Z_{2, J c}^{\mu^{\prime}}\right)$ for some $\mu^{\prime} \neq \mu$.

Calling $\mathbb{E}_{A}$ for $A=J$ or $A=J^{c}$ the expectation with respect to the random variables $\left(\xi_{i}^{\mu^{\prime}}\right)\left(\xi_{i}^{\mu}\right)$ for $i \in A$, and taking into account the previous remark, we get

$$
\begin{equation*}
\mathbb{E}\left(e^{i Z_{1, J}\left(Z_{2, J}+Z_{2, J c}\right)}\right)=\mathbb{E}_{J}\left(e^{i Z_{1, J} Z_{2, J}} \mathbb{E}_{J c}\left(e^{i Z_{1, J} Z_{2, J c}}\right)\right) \tag{2.9}
\end{equation*}
$$

Now looking at (2.5), it is easy to check that

$$
\begin{align*}
E_{J c}\left(e^{i Z_{1, J} Z_{2, j c}}\right)= & \sum_{n=0}^{\left|J^{c}\right|}\binom{\left|J^{c}\right|}{n}\left(\frac{1}{q}\right)^{n}\left(1-\frac{1}{q}\right)^{\left|J^{c \mid}\right|-n} \\
& \times \exp Z_{1, J} \frac{2 t}{\left(N V_{2}\right)^{1 / 2}}\left(n-\frac{\left|J^{c}\right|}{q}\right) \tag{2.10}
\end{align*}
$$

Now in order to perform the expectation with respect to $\mathbb{E}_{J}$ we need the explicit form of the joint distribution of $\left(Z_{1, J}, Z_{2, J}\right)$. Assume for the moment that for any function $f: \mathbb{R}^{2} \rightarrow \mathbb{R}$,

$$
\begin{align*}
& \mathbb{E}_{J}\left(f\left(Z_{2, J}, Z_{1, J}\right)\right) \\
&= \sum_{K=0}^{|J|} \sum_{l=-K}^{+K}\binom{J}{K}\binom{K}{(K+l) / 2}\left(1-\frac{2}{q}\right)^{|J|-K}\binom{2}{q}^{K}\left(\frac{1}{2}\right)^{K} \\
& \times f\left(\frac{K-2|J| / q}{\left(N V_{2}\right)^{1 / 2}}, \frac{l}{\left(N V_{1}\right)^{1 / 2}}\right) \tag{2.11}
\end{align*}
$$

Therefore we get

$$
\begin{align*}
& \mathbb{E}\left(e^{i Z_{1, j}\left(Z_{2,1}+Z_{2, J c}\right)}\right. \\
& =\sum_{K=0}^{|J|} \sum_{l=-K}^{+K} \sum_{n=0}^{\left|J^{c}\right|}\binom{|J|}{K}\binom{K}{(K+l) / 2}\binom{\left|J^{c}\right|}{n}\left(1-\frac{2}{q}\right)^{|J|-K}\left(\frac{2}{q}\right)^{K}\left(\frac{1}{2}\right)^{K} \\
& \times\left(1-\frac{1}{q}\right)^{\mid J^{C \mid}-n}\left(\frac{1}{q}\right)^{n} \exp \frac{t l}{\left(N V_{1}\right)^{1 / 2}}\left(\frac{(K-2 J / q)+2\left(n-\left|J^{c}\right| / q\right)}{\left(V_{2} N\right)^{1 / 2}}\right) \tag{2.12}
\end{align*}
$$

Let us now check (2.11).
For each $K \in 0, \ldots,|J|$, we can choose $A,|A|=|J|-K$, such that, for $i \in A$,

$$
\delta\left(\xi_{i}^{\mu^{\prime}},\left(\xi_{J}^{\mu}\right)_{i}\right)+\delta\left(\xi_{i}^{\mu^{\prime}}, \xi_{i}^{\mu}\right)=0
$$

and for $i \in J \backslash A$

$$
\delta\left(\xi_{i}^{\mu^{\prime}},\left(\xi_{J}^{\mu}\right)_{i}\right)+\delta\left(\xi_{i}^{\mu^{\prime}}, \xi_{i}^{\mu}\right)=1
$$

This choice can be done in $\binom{|J|}{K}$ different ways and the probability of any choice is $(1-2 / q)^{1 \cdot J}-K^{K}(2 / q)^{K}$.

Now for fixed $K$ and $A$, we choose on $J \backslash A$ a subset $B$ of cardinal $(K+l) / 2$, where for any $j \in B$,

$$
\delta\left(\xi_{i}^{\mu^{\prime}},\left(\xi_{J}^{\mu}\right)_{i}\right)-\delta\left(\zeta_{i}^{\mu^{\prime}}, \xi_{i}^{\mu}\right)=1
$$

and for any $j \in(Л A) \backslash B$,

$$
\delta\left(\xi_{i}^{\mu^{\prime}},\left(\xi_{J}^{\mu}\right)_{i}\right)-\delta\left(\xi_{i}^{\mu^{\prime}}, \xi_{i}^{\mu}\right)=-1
$$

We have $\left({ }_{(K+l / 2}^{K}\right)$ choices of such subset $B$ and each choice has probability $(1 / 2)^{K}$. Note that

$$
\sum_{i \in J \backslash A} \delta\left(\xi_{i}^{\mu^{\prime}},\left(\xi_{J}^{\mu}\right)_{i}\right)-\delta\left(\xi_{i}^{\mu^{\prime}}, \xi_{i}^{\mu}\right)=l
$$

In order to estimate (2.12), we will use the two following lemmas.
Lemma 2.1. For any real number $\lambda$ and any integer $K \geqslant 0$

$$
\begin{equation*}
\sum_{l=-K}^{+K}\binom{K}{(K+l) / 2}\left(\frac{1}{2}\right)^{k} e^{\lambda I} \leqslant e^{\lambda^{2} K / 2} \tag{2.13}
\end{equation*}
$$

Proof. This follows from

$$
(\cosh \lambda)^{K} \leqslant \exp \frac{\lambda^{2}}{2} K
$$

Lemma 2.2. Let $S_{n}=\sum_{i=1}^{n} X_{i}$, where $X_{i}=0$ with probability $p$ and $X_{i}=1$ with probability $1-p=r$. Then for any real number $\lambda$

$$
\begin{equation*}
\mathbb{E}\left(e^{\lambda\left(S_{n}-\mathbb{E}\left(S_{n}\right)\right)}\right) \leqslant e^{\left(\lambda^{2} / 2\right) n} \tag{2.14}
\end{equation*}
$$

Proof. Using the Jensen inequality, we get, for any $i \in\{1, \ldots, n\}$,

$$
\begin{equation*}
\mathbb{E}\left(e^{\lambda\left(X_{i}-\mathbb{E}\left(X_{i}\right)\right)}\right) \leqslant \mathbb{E}\left(e^{\lambda\left(X_{i}-X_{i}^{\prime}\right)}\right) \tag{2.15}
\end{equation*}
$$

when $X_{i}^{\prime}$ have the same distribution as $X_{i}$ and is independent of $X_{i}$. Remark that $X_{i}-X_{i}^{\prime}= \pm 1$ with probability $r p$ and $X_{i}-X_{i}^{\prime}=0$ with probability $p^{2}+r^{2}$. Therefore

$$
\mathbb{E}\left(e^{\lambda\left(X_{i}-X_{i}^{\prime}\right)}\right)=p^{2}+r^{2}+2 p r\left(\frac{e^{\lambda}+e^{-\hat{\lambda}}}{2}\right) \leqslant \cosh \lambda \leqslant \exp \frac{\lambda^{2}}{2}
$$

Now using Lemma 2.1, we get, for any $K \geqslant 0$,

$$
\begin{align*}
& \sum_{l=-K}^{+K}\binom{K}{(K+l) / 2}\left(\frac{1}{2}\right)^{K} \exp \frac{t l}{\left(N V_{1}\right)^{1 / 2}}\left(\frac{(K-2|J| / q)+2\left(n-\left|J^{c}\right| / q\right)}{\left(V_{2} N\right)^{1 / 2}}\right) \\
& \quad \leqslant \exp \frac{t^{2}}{2} \frac{K}{N V_{1}}\left(\frac{(K-2|J| / q)+2\left(n-\left|J^{c}\right| / q\right)}{\left(V_{2} N\right)^{1 / 2}}\right)^{2} \\
& \left.\quad \leqslant \mathbb{E}\left[\exp \left(\frac{t \sqrt{\delta}}{\sqrt{V_{1}}} Z\right)\right] \frac{(K-2|J| / q)+2\left(n-\left|J^{c}\right| / q\right.}{\left(V_{2} N\right)^{1 / 2}}\right) \tag{2.16}
\end{align*}
$$

where $Z$ is a normalized Gaussian random variable and we have used $K / N \leqslant \delta$.

Now inserting (2.16) in (2.12) and using Lemma 2.2 with $p_{1}=$ $(1-2 / q)$ and $p_{2}=(1-1 / q)$, we get

$$
\begin{align*}
\mathbb{E}\left(e^{t Z_{1, \jmath}\left(Z_{2, t}+Z_{2, s}\right)}\right) & \leqslant \mathbb{E}\left(\exp \frac{t^{2}}{2} \frac{\delta Z^{2}}{V_{1} V_{2}}\left(\frac{|J|+4\left|J^{c}\right|}{N}\right)\right) \\
& =\left(1-t^{2} \frac{\delta(\delta+4(1-\delta))}{V_{1} V_{2}}\right)^{-1 / 2} \tag{2.17}
\end{align*}
$$

Calling $v=\delta(\delta+4(1-\delta)) / V_{1} V_{2}$, we have

$$
\begin{equation*}
\inf _{t \geqslant 0} e^{-t u} \frac{1}{\left(1-t^{2} v\right)^{1 / 2}}=e^{\Gamma(u, v)} \tag{2.18}
\end{equation*}
$$

where

$$
\begin{equation*}
\Gamma(u, v)=\frac{1}{2}\left\{\left(1+\frac{4 u^{2}}{v}\right)^{1 / 2}-1+\log \frac{v}{u^{2}}\left[\left(1+\frac{4 u^{2}}{v}\right)^{1 / 2}-1\right]\right\} \tag{2.19}
\end{equation*}
$$

Therefore, using (2.19), (2.18), (2.8), and (2.2), if we prove that

$$
\begin{equation*}
-\delta \log \delta-(1-\delta) \log (1-\delta)+\delta \log (q-1)-\alpha \Gamma(u, v)<0 \tag{2.20}
\end{equation*}
$$

then using the Borel-Cantelli lemma, we can finish the proof of the theorem.

Recall

$$
u=\frac{\delta[(1-\delta)+(1-2 / q)]-\varepsilon}{\alpha\left[(2 \delta / q) V_{2}\right]^{1 / 2}}
$$

and $\lim _{\delta \rightarrow 0} V_{2}=4(1 / q)(1-1 / q)$. Choose for convenience $\varepsilon=\delta(1-2 / q)$; then

$$
u \sim \frac{\sqrt{\delta}}{\alpha} C_{1}(q) \xrightarrow[\alpha \rightarrow 0]{ } \infty
$$

for some constant $C_{1}(q)$.
Now, it can be checked that $\lim _{u \rightarrow \infty}[\Gamma(u, v) / u]=1 / \sqrt{v}$; therefore $\alpha \Gamma(u, v) \sim \sqrt{\delta} C_{2}(q)$ for some constant $C_{2}(q)$.

Now

$$
\begin{aligned}
& \frac{C_{2}(q) \sqrt{\delta}}{-\delta \log \delta-(1-\delta) \log (1-\delta)+\delta \log (q-1)} \\
& \sim \frac{C_{2}(q)}{-\sqrt{\delta} \log \delta+\sqrt{\delta} \log (q-1)} \xrightarrow[\delta \rightarrow 0]{ }+\infty
\end{aligned}
$$

This proves (2.20), choosing first $\alpha$ small enough and then $\delta$.

Remark. Making an explicit computation, if we are interested in $q$ very large, it can be checked that as $q \rightarrow \infty, C_{1}(q) \sim\left[(q \delta)^{1 / 2} / \alpha\right] C_{1}$, $C_{2}(q) \sim(q / \delta)^{1 / 2}\left(C_{2}\right.$ and therefore

$$
\begin{array}{r}
\frac{(\delta / q)^{1 / 2}}{-\delta \log \delta-(1-\delta) \log (1-\delta)+\delta \log (q-1)} \\
\sim \frac{\sqrt{q}}{-\sqrt{\delta} \log \delta+\sqrt{\delta} \log (q-1)} \rightarrow+\infty
\end{array}
$$

by choosing an appropriate $\delta$ depending on $q$.
At this point we remark that in the limit $q \rightarrow \infty$ the estimate (2.14) is too poor to take seriously when one considers what happens when $q \rightarrow \infty$.

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